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ON THE NEGATIVE BINOMIAL CONVERGENCE IN A CLASS OF M-DIMENSIONAL ETC(U)
JUL 78 H LACAYO, N A LANGBERG AFOSR-74-2581

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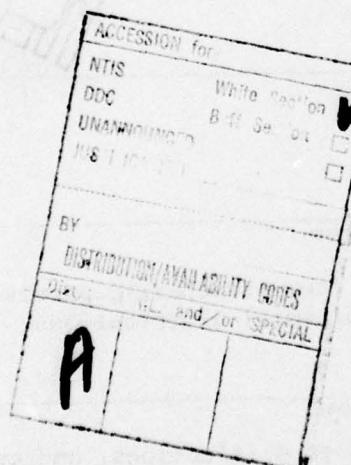
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20. Abstract

$\alpha_i X_{N,i}(t) [1 - \frac{1}{N} \sum_{i=1}^m (X_{N,i}(t) - X_{N,i}(0))]$, $i = 1, \dots, m$, $\alpha_1, \dots, \alpha_m > 0$ and

if $\lim_{N \rightarrow \infty} X_{N,i}(0) = b_i \in \{1, 2, \dots\}$, then $(X_{N,1}(t), \dots, X_{N,m}(t))$ converges as
 $N \rightarrow \infty$ to a random vector with independent negative binomial components.



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On the Negative Binomial Convergence
in a Class of m -Dimensional Simple Epidemics

by

H. Lacayo and Naftali A. Langberg

ABSTRACT

We consider a population which is exposed to m infections, and consist initially of N susceptibles. At each point in time at most one susceptible becomes infective, and only from one cause. This m -dimensional simple epidemic is a stochastic process, $(X_{N,1}(t), \dots, X_{N,m}(t))$, with components counting the number of infectives from the respective causes at time t . We show that if the transition rates of cause 1 through m at time t are given by $\alpha_i X_{N,i}(t)[1 - \frac{1}{N} \sum_{i=1}^m (X_{N,i}(t) - X_{N,i}(0))]$ $i = 1, \dots, m$, $\alpha_1, \dots, \alpha_m > 0$ and if $\lim_{N \rightarrow \infty} X_{N,i}(0) = b_i \in \{1, 2, \dots\}$, then $(X_{N,1}(t), \dots, X_{N,m}(t))$ converges as $N \rightarrow \infty$ to a random vector with independent negative binomial components.

Key Words: Exponential distributions, convergence in law, stochastic processes, m -dimensional simple epidemics, negative binomial distributions.

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1. Introduction and Summary

Introduction.

We consider a population which is exposed to m infections and consists initially of N susceptibles and $b_{N,i}$ infectives from cause i , $i = 1, \dots, m$. At each point in time at most one susceptible becomes an infective, and only from one cause. Once an individual enters the infective state he remains there and cannot be a carrier of any other infection. This m -dimensional simple epidemic is a stochastic process, $(X_{N,1}(t), \dots, X_{N,m}(t))$, with components counting the number of infectives from the respective causes at time t . We show that if the transition rates of cause 1 through m at time t are given by:

$$\alpha_i X_{N,i}(t) [1 - \frac{1}{N} \sum_{i=1}^m (X_{N,i}(t) - b_{N,i})], \quad i = 1, \dots, m, \quad \alpha_1, \dots, \alpha_m > 0, \quad (1.1)$$

and if

$$\lim_{N \rightarrow \infty} b_{N,i} = b_i \in \{1, 2, \dots\}, \quad i = 1, \dots, m \quad (1.2)$$

then for every positive real numbers t and β , and m nonnegative integers k_1, \dots, k_m ,

$$\lim_{N \rightarrow \infty} P\{X_{N,i}(t) \geq k_i + b_i, i = 1, \dots, m\} = \prod_{i=1}^m P\{X_i(t) \geq k_i + b_i\} \quad (1.3)$$

and

$$\lim_{N \rightarrow \infty} E(X_{N,i}(t))^\beta = E(X_i(t))^\beta, \quad i = 1, \dots, m, \quad (1.4)$$

where $X_1(t), \dots, X_m(t)$ are negative binomial random variables with

respective parameters b_i and $e^{-\alpha_i t}$, $i = 1, \dots, m$. The only previous work in the area of m -dimensional simple epidemics is the one by Billard, Lacayo and Langberg (1978), who proved (1.3) and (1.4) with the additional assumption that $\alpha_1 = \alpha_2 \dots = \alpha_m$. In that case the interinfection random times are independent of the infection causes. Generally, as will be pointed out in Section 2 this property does not hold. The proof of (1.3) in the cited reference depends on the special structure of the interinfection times. To obtain (1.3) and (1.4) we have approached the problem from an alternate viewpoint.

Summary.

In Section 2 we present a rigorous definition of an m -dimensional simple epidemic, and describe the ones that are the subject of our analysis. Statements (1.3) and (1.4) are proven in Section 3. In this section we present an m -dimensional simple epidemic in a population consisting initially of infinitely many susceptibles. This theoretical simple epidemic is instrumental in the proofs of (1.3) and (1.4).

2. The m -dimensional Simple Epidemic

An m -dimensional simple epidemic in a population consisting initially of N susceptibles and $b_{N,i}$ infectives from cause i , $i = 1, \dots, m$ can be described by N bivariate random vectors $(T_{N,1}, \xi_{N,1}), \dots, (T_{N,N}, \xi_{N,N})$. $T_{N,1}, \dots, T_{N,N}$ are the random interinfection times, and $\xi_{N,1}, \dots, \xi_{N,N}$ are discrete random variables with values in $\{1, \dots, m\}$ designating the respective infection causes. Let $S_{N,-}$ be equal to $\sum_{j=1}^k T_{N,j}$, $k = 1, \dots, N$. The finite dimensional joint distributions of $(X_{N,1}(t), \dots, X_{N,m}(t))$ are determined by $(T_{N,1}, \xi_{N,1}), \dots, (T_{N,N}, \xi_{N,N})$, through the following set equalities:

$$X_{N,i}(t) = b_{N,i} + k = \sum_{r=k}^N (S_{N,r} \leq t < S_{N,r+1}, \sum_{q=1}^r I(\xi_{N,q} = 1) = k), \quad t > 0 \quad (2.1)$$

$k = 0, \dots, N$, and $i = 1, \dots, m$. ($S_{N,0} = 0, S_{N,N+1} = \infty$).

We assume throughout that $(T_{N,1}, \xi_{N,1}), \dots, (T_{N,N}, \xi_{N,N})$, are determined by (2.2) and (2.3).

$$P\{\xi_{N,j} = \iota_j, j = 1, \dots, N\} = \\ \left[\prod_{i=1}^m \alpha_i \prod_{j=0}^{r_i-1} (b_{N,i} + j) \right] \left[\prod_{k=1}^N \left(\sum_{i=1}^m \alpha_i (b_{N,i} + \sum_{q=1}^{k-1} I(\iota_q = i)) \right)^{-1} \right], \quad (2.2)$$

where $r_i = \sum_{j=1}^N I(\iota_j = i)$, $i = 1, \dots, m$.

$$P\{T_{N,k} | \xi_{N,1}, \dots, \xi_{N,k-1} > t_k, k = 1, \dots, N\} = e^{-\sum_{j=1}^N \mu_j t_j} \quad (2.3)$$

where t_1, \dots, t_n are positive real numbers, and μ_j is equal to $\frac{N-j+1}{N} \sum_{i=1}^m \alpha_i (b_i + \sum_{q=1}^{j-1} I(\iota_q = i))$, for $j = 1, \dots, N$.

In the particular case discussed in Billard, Lacayo and Langberg (1978), condition (2.3) reduces to independent exponential interinfection times

with rates equal to $\alpha \frac{N - k + 1}{N} (\sum_{i=1}^m b_{N,i} + k - 1)$, $k = 1, \dots, N$, independent of $\xi_{N,1}, \dots, \xi_{N,N}$. In addition equation (2.2) simplifies to $[\prod_{i=1}^m \prod_{j=0}^{k-1} (b_{N,i} + j)] [\prod_{k=1}^N (\sum_{i=1}^m b_{N,i} + k - 1)]^{-1}$. Finally we note that if (2.2) and (2.3) hold, then the m -dimensional simple epidemic satisfies (1.1). Conversely, if (1.1) is satisfied, then equations (2.2) and (2.3) follow by the inherited Markovian structure of the m -dimensional simple epidemic.

3. Main Results

Let $(X_1(t), \dots, X_m(t))$ be an m -dimensional simple epidemic in a population consisting initially of b_i infectives from cause i , $i = 1, 2, \dots, m$, and infinitely many susceptibles. This epidemic can be described by a sequence of bivariate random vectors, $(T_1, \xi_1), (T_2, \xi_2), \dots, (T_k, \xi_k)$. T_1, T_2, \dots are the interinfection times, and ξ_1, ξ_2, \dots are random variables with values in $\{1, \dots, m\}$, which designate the respective infection causes. Let S_k be equal to $\sum_{j=1}^k T_j$, $k = 1, 2, \dots$. The finite dimensional joint distributions of $(X_1(t), X_m(t))$ are determined by (T_k, ξ_k) , $k = 1, 2, \dots$, through the following set equalities:

$$(X_i(t) = b_i + k) = \bigcup_{r=k}^{\infty} (S_r \leq t < S_{r+1}, \sum_{q=1}^r I(\xi_q = i) = k) \quad (3.1)$$

for every positive real number t , positive integer k , and $i = 1, \dots, m$.

We assume throughout that the sequence $(T_1, \xi_1), (T_2, \xi_2), \dots$, is determined by (3.2) and (3.3).

$$\begin{aligned} P\{\xi_k = k, k = 1, \dots, n\} = \\ \left[\prod_{i=1}^m \prod_{j=0}^{r_i-1} (b_i + j) \right] \left[\prod_{j=1}^n \left(\sum_{i=1}^m \alpha_i (b_i + \sum_{q=1}^{j-1} I(\xi_q = 1)) \right)^{-1} \right] \end{aligned} \quad (3.2)$$

where $r_i = \sum_{q=1}^n I(\xi_q = i)$, $i = 1, \dots, m$, and $n = 1, 2, \dots$.

$$P\{T_k | \xi_1, \dots, \xi_{k-1} > t_k, k = 1, \dots, n\} = e^{-\sum_{j=1}^n \mu_j t_j} \quad (3.3)$$

where t_1, \dots, t_n are positive real numbers, $\mu_j = \sum_{i=1}^m \alpha_i (b_i + \sum_{q=1}^{j-1} I(\xi_q = 1))$, $j = 1, \dots, n$, and $n = 1, 2, \dots$.

We will show that for every two positive real numbers t and β

$(X_{N,1}(t), \dots, X_{N,m}(t))$ converges in law to $(X_1(t), \dots, X_m(t))$ as $N \rightarrow \infty$, and that $\lim_{N \rightarrow \infty} E(X_{N,i}(t))^\beta = E(X_i(t))^\beta$ for $i = 1, \dots, m$. First we introduce two lemmas.

Lemma 3.1. Let U_1, U_2, \dots be an i.i.d. sequence of exponential random variables with mean 1. Further let $(\tau_{b+r-1,1}, \dots, \tau_{b+r-1,b+r-1})$ be the order statistic of a sample of size $b + r - 1$ from U_1 . Then $\sum_{j=1}^{r-1} (b + j - 1)^{-1} U_j$ and $\tau_{b+r-1,r}$ are identically distributed.

Proof. It suffices to note that the spacings; $\tau_{r+b-1,r-j+1} - \tau_{r+j-1,r-j}$, $j = 1, \dots, r$, ($\tau_{r+b-1,0} = 0$) are independent exponential random variables with means equal respectively to $(b + j - 1)^{-1}$, $j = 1, \dots, r$.

Lemma 3.2. Let t and β be positive real numbers; let $c = \sup_N \sum_{i=1}^m b_{N,i}$ and $\alpha = \max_{1 \leq i \leq m} \alpha_i$. If (1.2) holds then; (i) $P(S_{N,r} \leq t) \leq (1 - e^{-act})^r$ for $r = 1, \dots, N$, and (ii) $\sup_N E(X_{N,i}(t))^\beta < \infty$, for $i = 1, \dots, m$.

Proof. We note that (ii) is a corollary of (i). To prove (i) let U_1, U_2, \dots be as in Lemma 3.1. Since $P(S_{N,k} \leq t) \leq P\{\sum_{j=1}^r j^{-1} U_j \leq act\}$, (i) follows from Lemma 3.1.

Theorem 3.3. Let β and t be positive real numbers. If (1.2) holds then;

(i) $(X_{N,1}(t), \dots, X_{N,m}(t))$ converges in law as $N \rightarrow \infty$ to $(X_1(t), \dots, X_m(t))$, and (ii) $\lim_{N \rightarrow \infty} E(X_{N,i}(t))^\beta = E(X_i(t))^\beta$ for $i = 1, \dots, m$.

Proof. We note that (ii) is a corollary of (i) and Lemma 3.2 (ii).

To prove (i) let k_1, \dots, k_m be m nonnegative integers, A be the set $\{(r_1, \dots, r_m) | r_i \geq k_i, i = 1, \dots, m\}$ and let $R = \sum_{i=1}^m r_i$. Since for N sufficiently large, $P\{X_{N,i}(t) \geq k_i + b_i, i = 1, \dots, m\} = \sum_A P\{S_{N,R} \leq t < S_{N,R+1}, \sum_{q=1}^R I(\xi_{N,q} = i) = r_i, i = 1, \dots, m\} I(R \leq N\}$, (i) follows from Lemma (3.2) (i), the dominated convergence theorem, and some simple calculations.

For reference purposes the following two well known results are presented.

Lemma 3.4. Let U_1, U_2, \dots be as in Lemma 3.1, and let μ_1, μ_2, \dots be a sequence of positive real numbers. Further we denote by f_k the density function of $\sum_{j=1}^k \mu_j^{-1} U_j$, $k = 1, 2, \dots$. Then, (i) $\mu_{r+1}^{-1} f_{r+1}(t) = P\{\sum_{j=1}^r \mu_j^{-1} U_j \leq t < \sum_{j=1}^{r+1} \mu_j^{-1} U_j\}$, and (ii) $f_r(t) = (\prod_{j=1}^r \mu_j) \sum_{j=1}^r e^{-\mu_j t} \prod_{\substack{i=1 \\ i \neq j}}^r (\mu_i - \mu_j)^{-1}$, provided μ_1, \dots, μ_r are distinct.

We are ready to state and prove the main result of the paper.

Theorem 3.5. Let t and β be positive real numbers, and k_1, \dots, k_m be nonnegative integers. Further let A be the set $\{r_1, \dots, r_m | r_i \geq k_i, i = 1, \dots, m\}$. If conditions (1.1) and (1.2) are satisfied, then the following hold.

- (i) $\lim_{N \rightarrow \infty} P\{X_{N,i}(t) \geq k_i + b_i, i = 1, \dots, m\} = \sum_A \prod_{i=1}^m \begin{pmatrix} b_i + r_i - 1 \\ b_i - 1 \end{pmatrix} e^{-\alpha_i b_i t} (1 - e^{-\alpha_i t})^{k_i}$.
- (ii) $\lim_{N \rightarrow \infty} E(X_{N,i}(t))^\beta$ is the β th moment of a negative binomial random variable with parameters b_i and $e^{-\alpha_i t}$, $i = 1, \dots, m$.

Proof. By Theorem 3.3 it suffices to show that $X_1(t), \dots, X_m(t)$ are independent, and have the respective negative binomial marginal distributions.

Let k_1, \dots, k_m be nonnegative integers, with sum equal to k . Further, let B denote the set

$$[(\ell_1, \dots, \ell_k) | \ell_1, \dots, \ell_k \in \{1, \dots, m\}, \sum_{q=1}^k I(\ell_q = i) = k_i, i = 1, \dots, m].$$

We note that $P\{X_i(t) = k_i + b_i, i = 1, \dots, m\}$ equals

$$\sum_B P\{S_k \leq t < S_{k+1} | \xi_q = \ell_q, q = 1, \dots, k\} P\{\xi_q = \ell_q, q = 1, \dots, k\}.$$

From Lemma 3.4 and equation (3.2) it follows that $P\{X_i(t) = k_i + b_i, i = 1, \dots, m\}$ equals

$$\prod_{i=1}^m \binom{b_i + k_i - 1}{b_i - 1} e^{-a_i b_i t} \alpha_i^{k_i} k_i! \sum_B \sum_{j=1}^{k+1} e^{-\theta_j t} \prod_{\substack{i=1 \\ i \neq j}}^{k+1} (\theta_i - \theta_j)^{-1}, \text{ where}$$

$$\theta_j = \sum_{q=1}^{j-1} \sum_{i=1}^m \alpha_i I(\ell_q = i), j = 1, \dots, k+1.$$

By reusing Lemma 3.4 we obtain that $P\{X_i(t) = k_i + b_i, i = 1, \dots, k\}$ equals

$$\prod_{i=1}^m \binom{b_i + k_i - 1}{b_i - 1} e^{-a_i b_i t} \alpha_i^{k_i} k_i! \sum_B P\left\{ \sum_{j=1}^{k+1} \theta_j^{-1} U_j \leq t \right\} \prod_{j=2}^{k+1} \theta_j^{-1}.$$

To complete the proof it suffices to show that

$$\prod_{i=1}^m \alpha_i^{k_i} k_i! \sum_B P\left\{ \sum_{j=2}^{k+1} \theta_j^{-1} U_j \leq t \right\} \prod_{j=2}^{k+1} \theta_j^{-1} = \prod_{i=1}^m (1 - e^{-a_i t})^{k_i}. \quad (3.4)$$

Let F denote the left side of equation (3.4). Further let

$$B_i = [(\ell_1, \ell_2, \dots, \ell_k) | \ell_2, \dots, \ell_k \in \{1, \dots, m\}, \sum_{q=2}^k I(\ell_q = r) = k_r, r = 1, \dots, m, r \neq i, \sum_{q=2}^k I(\ell_q = i) = k_i - 1].$$

We observe that $F(t)$ equals $\prod_{i=1}^m \alpha_i^{k_i} k_i!$ $\sum_{i=1}^m \sum_{B_i} P\left\{\sum_{j=2}^{k+1} \theta_j^{-1} U_j \leq t\right\} \prod_{j=2}^{k+1} \theta_j^{-1}$.

By a simple calculation and Lemma 3.4, it follows that

$$\frac{dF}{dt} = \prod_{i=1}^m \alpha_i^{k_i} k_i! \sum_{i=1}^m e^{-\alpha_i t} \sum_{B_i} P\left\{\sum_{j=3}^{k+1} \theta_j^{-1} U_j \leq t\right\} \prod_{j=3}^{k+1} \theta_j^{-1}, \text{ where } \theta_{j,i} = \sum_{r=1}^m \alpha_r \sum_{q=2}^{j-1} I(\ell_q = r), i = 1, \dots, m, j = 3, \dots, k+1. \text{ Consequently Equation}$$

$$\sum_{i=1}^m \alpha_i \sum_{B_i} P\left\{\sum_{j=3}^{k+2} \theta_j^{-1} U_j \leq t\right\} \prod_{j=3}^{k+1} \theta_j^{-1}, i = 1, \dots, m.$$

(3.4) is established by an induction argument on k applied to each of the

$$\sum_{B_i} P\left\{\sum_{j=3}^{k+2} \theta_j^{-1} U_j \leq t\right\} \prod_{j=3}^{k+1} \theta_j^{-1}, i = 1, \dots, m.$$

Corollary 3.6. Let $X_N(t)$ the total number of infectives at time t . If

(1.1) and (1.2) are satisfied, then the following hold.

- (i) $X_N(t)$ converges in law as $N \rightarrow \infty$ to a sum of m independent negative binomial random variables with respective parameters $b_i, e^{-\alpha_i t}, i = 1, \dots, m$, and
- (ii) $\lim_{N \rightarrow \infty} E(X_N(t))^\beta$ equals to the β th moment of the limiting random variable given in (i).

In the particular case discussed in Billard, Lacayo and Langberg (1978),

$X_N(t)$ converges in law to a negative binomial random variable with parameters

$\sum_{i=1}^m b_i$ and e^{-at} . This fact played a major role in the proof of (1.3) for the symmetric case as presented in the cited reference. Finally we note

that it is of interest to investigate the asymptotic behavior of m -dimensional simple epidemics, that do not have the Markovian structure.

REFERENCE

- (1) Billard, L., Lacayo, H. and Langberg, Naftali A. (1978). The Symmetric
m-dimensional Simple Epidemic. FSU Statistics Report M461.

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